



Economics and Finance: *q***-Statistical Stylized Features Galore**

Constantino Tsallis 1,2,3

Review

- ¹ Centro Brasileiro de Pesquisas Físicas and National Institute of Science and Technology for Complex Systems, Rua Xavier Sigaud 150, Urca, Rio de Janeiro 22290-180, Brazil; tsallis@cbpf.br; Tel.: +55-(21)-2141-7190
- ² Santa Fe Institute, 1399 Hyde Park Road, Santa Fe, NM 87501, USA
- ³ Complexity Science Hub Vienna, Josefstädter Strasse 39, Vienna 1080, Austria

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Abstract: The Boltzmann–Gibbs (BG) entropy and its associated statistical mechanics were generalized, three decades ago, on the basis of the nonadditive entropy S_q ($q \in \mathcal{R}$), which recovers the BG entropy in the $q \rightarrow 1$ limit. The optimization of S_q under appropriate simple constraints straightforwardly yields the so-called *q*-exponential and *q*-Gaussian distributions, respectively generalizing the exponential and Gaussian ones, recovered for q = 1. These generalized functions ubiquitously emerge in complex systems, especially as economic and financial stylized features. These include price returns and volumes distributions, inter-occurrence times, characterization of wealth distributions and associated inequalities, among others. Here, we briefly review the basic concepts of this *q*-statistical generalization and focus on its rapidly growing applications in economics and finance.

Keywords: economics and finance; nonadditive entropies; nonextensive statistical mechanics

1. Introduction

Exponential and Gaussian functions ubiquitously emerge within linear theories in mathematics, physics, economics and elsewhere. To illustrate in what sense they are linear, let us focus on three typical mathematical situations, namely an ordinary differential equation, a partial derivative equation and an entropic optimization.

Consider the following ordinary differential equation:

$$\frac{dy}{dx} = ay \quad [y(0) = 1]. \tag{1}$$

The solution is the well-known exponential function:

$$y = e^{ax} \,. \tag{2}$$

Consider now the following partial derivative equation:

$$\frac{\partial p(x,t)}{\partial t} = D \frac{\partial^2 p(x,t)}{\partial x^2} \quad [D > 0; t \ge 0; p(x,0) = \delta(x)],$$
(3)

where $\delta(x)$ is the Dirac delta function. The solution is the well-known Gaussian distribution:

$$p(x,t) = \frac{1}{\sqrt{2\pi Dt}} e^{-x^2/2Dt} \,. \tag{4}$$

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Let us finally consider the following entropic functional:

$$S_{BG} = -k \int dx \, p(x) \ln p(x) \quad (k > 0) \tag{5}$$

with the constraint:

$$\int dx \, p(x) = 1 \,, \tag{6}$$

where BG stands for Boltzmann–Gibbs; *k* is a conventional positive constant (usually $k = k_B$ in physics, and k = 1 elsewhere). If we optimize the functional (5) with the constraint (6) and:

$$\langle \epsilon(x) \rangle \equiv \int dx \, p(x) \epsilon(x) = u \quad (u \in \mathcal{R})$$
 (7)

 $\epsilon(x)$ being bounded below, we obtain:

$$p(x) = \frac{e^{-\beta\epsilon(x)}}{\int dx \, e^{-\beta\epsilon(x)}},\tag{8}$$

where $\beta \equiv 1/kT$ is the Lagrange parameter associated with Constraint (7); *T* is the absolute temperature within BG statistical mechanics (necessarily T > 0 if $\epsilon(x)$ is unbounded from above; but both T > 0 and T < 0 possibilities exist if $\epsilon(x)$ is bounded also from above). The probability distribution p(x) corresponds to the celebrated BG weight, where $Z \equiv \int dx e^{-\beta\epsilon(x)}$ is usually referred to as the partition function. Two particular cases emerge frequently. The first of them is $\epsilon(x) = x$ ($x \ge 0$) with $u \equiv \langle x \rangle$, hence $p(x) = \frac{e^{-x/\langle x \rangle}}{\langle x \rangle}$, thus recovering solution (2). The second one is $\epsilon(x) = x^2$ with $u \equiv \langle x^2 \rangle$, hence $p(x) = \frac{e^{-x^2/2\langle x^2 \rangle}}{\sqrt{2\pi\langle x^2 \rangle}}$, thus recovering solution (4). Therefore, basic cases connect S_{BG} with the solutions of the linear Equations (1) and (3). In addition to that, let us make explicit in what sense S_{BG} is itself linear. We consider a system (A + B) constituted by two probabilistically independent subsystems A and B. In other words, we consider the case where the joint probability of (A + B) factorizes, i.e., $p^{(A+B)}(x,y) = p^{(A)}(x)p^{(B)}(y)$ [$\forall(x,y)$]. We straightforwardly verify that the functional S_{BG} is additive in the sense of Penrose [1], namely that:

$$S_{BG}(A+B) = S_{BG}(A) + S_{BG}(B).$$
(9)

In the present brief review, we shall address a special class of nonlinearities, namely those emerging within nonextensive statistical mechanics, *q*-statistics for short [2–6].

Equation (1) is now generalized into the following nonlinear one:

$$\frac{dy}{dx} = ay^q \quad [y(0) = 1; q \in \mathcal{R}].$$
(10)

Its solution is:

$$y = e_q^{ax} , \tag{11}$$

where the *q*-exponential function is defined as $e_q^z \equiv [1 + (1 - q)z]_+^{1/(1-q)}$ ($e_1^z = e^z$), with $[1 + (1 - q)z]_+ = 1 + (1 - q)z$ if z > 0 and zero otherwise. Its inverse function is the *q*-logarithm, defined as $\ln_q z \equiv \frac{z^{1-q}-1}{1-q}$ ($\ln_1 z = \ln z$). To avoid any confusion, let us mention that many other *q*-deformations of the exponential and logarithmic functions have been introduced in the literature for a variety of purposes; among them, we have for instance Ramanujan's *q*-exponential function, unrelated to the present one.

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Equation (3) is now generalized into the following nonlinear one (referred to in the literature as the porous medium equation [7–9]):

$$\frac{\partial p(x,t)}{\partial t} = D_q \frac{\partial^2 [p(x,t)]^{2-q}}{\partial x^2} \quad [D_q(2-q) > 0; D_1 \equiv D; q < 3; t \ge 0; p(x,0) = \delta(x)].$$
(12)

Its solution generalizes Equation (4) and is given by:

$$p(x,t) = \frac{1}{\sqrt{\pi A_q}} e_q^{-x^2/\{A_q[(D_q)t]^{\frac{2}{3-q}}\}}$$
(13)

with:

$$A_{q} = \begin{cases} \frac{\sqrt{q-1} \Gamma\left(\frac{1}{q-1}\right)}{\Gamma\left(\frac{3-q}{2(q-1)}\right)} & \text{if } 1 < q < 3, \\ 2 & \text{if } q = 1, \\ \frac{\sqrt{1-q} \Gamma\left(\frac{5-3q}{2(1-q)}\right)}{\Gamma\left(\frac{2-q}{1-q}\right)} & \text{if } q < 1. \end{cases}$$
(14)

Before going on, let us mention that solution (13) implies that x^2 scales like $t^{\frac{2}{3-q}}$, hence normal diffusion for q = 1, anomalous sub-diffusion for q < 1 and super-diffusion for 1 < q < 3, which has recently been impressively validated (within a 2% experimental error) in a granular medium [10]. The important connection between the power-law nonlinear diffusion (12) and the entropy S_q described here below was first established by Plastino and Plastino in [11], where they considered a more general evolution equation that reduces to (12) in the particular case of vanishing drift (i.e., F(x) = 0, $\forall x$). The Plastino–Plastino Equation [11] $\frac{\partial p(x,t)}{\partial t} = -\frac{\partial}{\partial x}[F(x)p(x,t)] + D_q \frac{\partial^2[p(x,t)]^{2-q}}{\partial x^2}$ with F(x) = -dV(x)/dx generalizes the porous medium equation in the same sense that the linear Fokker–Planck equation generalizes the classical heat equation. The above nonlinear Equations (10) and (12) have been addressed here in order to provide some basic mathematical structure to approaches of various economic- and financial-specific features presented later on.

Let us now focus on the entropic functional S_q upon which nonextensive statistical mechanics is based. It is defined as follows:

$$S_q \equiv k \frac{1 - \int dx \, [p(x)]^q}{q - 1} = k \int dx \, p(x) \ln_q \frac{1}{p(x)} = -k \int dx \, [p(x)]^q \ln_q p(x) = -k \int dx \, p(x) \ln_{2-q} p(x)$$
(15)

with $S_1 = S_{BG}$. If we optimize this functional with the constraints (6) and:

$$\langle \epsilon(x) \rangle_q \equiv \frac{\int dx \, [p(x)]^q \epsilon(x)}{\int dx \, [p(x)]^q} = u_q \quad (u_q \in \mathcal{R}; \, u_1 = u) \tag{16}$$

we obtain [4]:

$$p(x) = \frac{e_q^{-\beta_q \epsilon(x)}}{\int dx \, e_q^{-\beta_q \epsilon(x)}} = \frac{e_q^{-\beta'_q [\epsilon(x) - u_q]}}{\int dx \, e_q^{-\beta'_q [\epsilon(x) - u_q]}} \quad (\beta_1 = \beta'_1 = \beta) \,. \tag{17}$$

As before, two particular cases emerge frequently. The first of them is $\epsilon(x) = x$ ($x \ge 0$) with $u_q \equiv \langle x \rangle_q$; hence, p(x) recovers the form of (11). The second one is $\epsilon(x) = x^2$ with $u_q \equiv \langle x^2 \rangle_q$; hence, p(x) recovers the form of solution (13). Finally, if we consider S_q itself for two independent subsystems A and B, we straightforwardly verify the following nonlinear composition law:

$$\frac{S_q(A+B)}{k} = \frac{S_q(A)}{k} + \frac{S_q(B)}{k} + (1-q)\frac{S_q(A)}{k}\frac{S_q(B)}{k},$$
(18)

hence

$$S_q(A+B) = S_q(A) + S_q(B) + \frac{1-q}{k} S_q(A) S_q(B).$$
(19)

We then say that S_q is nonadditive for $q \neq 1$. Entropic additivity is recovered if $(1 - q)/k \rightarrow 0$, which can occur in two different circumstances: $q \rightarrow 1$ for fixed k or $k \rightarrow \infty$ for fixed q. Since k always appears in physics in the form kT, the limit $k \rightarrow \infty$ is equivalent to $T \rightarrow \infty$. This is, by the way, the basic reason for which, in the limit of high temperatures or low energies, Maxwell–Boltzmann statistics, Fermi–Dirac, Bose–Einstein and q-statistics asymptotically coincide.

The above *q*-generalized thermostatistical theory has been useful in the study of a considerable number of natural, artificial and social systems (see [12]). Theoretical and experimental illustrations in natural systems include long-range-interacting many-body classical Hamiltonian systems [13–20] (see also [21,22]; the study of the long-range version of [23] would surely be interesting), dissipative many-body systems [24], low-dimensional dissipative and conservative nonlinear dynamical systems [25–31], cold atoms [32–34], plasmas [35,36], trapped atoms [37], spin-glasses[38], power-law anomalous diffusion [39,40], granular matter [10], high-energy particle collisions [41–46], black holes and cosmology [47,48], chemistry [49], earthquakes [50], biology [51,52], solar wind [53,54], anomalous diffusion in relation to central limit theorems and overdamped systems [55–64], quantum entangled systems [65,66], quantum chaos [67], astronomical systems [68,69], thermal conductance [70], mathematical structures [71–76] and nonlinear quantum mechanics [77–96], among others. Illustrations in artificial systems include signal and image processing [97,98] and (asymptotically) scale-free networks [99–101]. In the realm of social systems, from now on, we focus on economics and financial theory [102–118].

2. Applications in Economics and Finance

2.1. Prices and Volumes

Time series of prices p_t (say of stocks, commodities, etc.), where t runs along chosen units (say seconds or minutes, or days, or years) are conveniently replaced by their corresponding returns (or logarithmic returns), defined as follows:

$$r_t = \ln \frac{p_{t+1}}{p_t} \simeq \frac{p_{t+1} - p_t}{p_t} \quad (t = 0, 1, 2, \dots) \,.$$
⁽²⁰⁾

Returns do not depend on the specific currency of the prices and fluctuate around zero; in addition to that, their definition cancels systematic inflation. The distribution of returns usefully characterizes the price fluctuations. See an illustration in Figure 1, from [103] (see also [104,118]). The amounts of the corresponding transactions are currently referred to as volumes: see, for example, Figure 2.

2.2. Volatilities

The volatility characterizes the size (standard deviation) of the fluctuations of returns. The volatility smile characterizes the correction of empirical volatilities with regard to a Gaussian-based expectation: see an illustration in Figure 3 (from [103]). To be more explicit, let us assume that we are handling the following Gaussian distribution $\propto e^{-Br^2}$, where \mathcal{B} characterizes univocally the volatility. To discuss the probability distribution of quantities such as \mathcal{B} , Queiros introduced [108] the *q*-log normal probability function:

$$p_q(x) = \frac{1}{Z_q x^q} e^{-\frac{|\ln q \ x - \mu|^2}{2\sigma^2}} \ (x > 0),$$
(21)

where Z_q is a normalizing factor and (q, μ, σ) are parameters. The q = 1 particular case corresponds to the standard log-normal function. See Figure 4 for illustrative examples of this function. See also Figure 5 for a real financial application.



Figure 1. Empirical return densities (points) and *q*-Gaussians (solid lines) for normalized returns of the 10 top-volume stocks in the NYSE and in NASDAQ in 2001. The dotted line is a (visibly inadequate) Gaussian distribution. The 2- and 3-min curves are moved vertically for display purposes. From [103]. There exist in the literature quite a few other such examples, for other stocks and other years, with similar values of *q*.



Figure 2. Empirical distributions (points) and *q*-exponential-like curves (solid lines) for normalized volumes of the 10 top-volume stocks in the NYSE and in the NASDAQ in 2001. The solid lines are fittings with a *q*-exponential multiplied by a power-law (analogous to the density of states prefactor that typically emerges for the distributions of quasi-particles in, say, condensed matter physics); from [103]. There exist in the literature quite a few other such examples, for other stocks and other years, with similar values of *q* and of the rest of the fitting indices.



Figure 3. Implied volatilities as a function of the strike price for call options on JY currency futures, traded on 16 May 2002, with 147 days left to expiration. In this typical example, the current price of a contract on Japanese futures is \$79, and the risk-free rate of return is 5.5%. Circles correspond to volatilities implied by the market, whereas triangles correspond to volatilities implied by our model with q = 1.4 and $\sigma = 10.2$ %. The dotted line is a guide to the eye. From [103].



Figure 4. Illustrations of the *q*-log-normal density for $\mu = 0$ and $\sigma = 1$: blue q = 5/4, red q = 1 and green q = 4/5. From [108].



Figure 5. Probability density function of a five-day volatility vs. \mathcal{B} . The symbols are obtained from the data, and the lines are the best fits with the Gamma distribution (dashed green) and the double-sided *q*-log-normal (red) with $\mu = 0.391$, $\sigma = 1.15$ and q = 1.22. For further details, see [108].

2.3. Inter-Occurrence Times

We can see in Figure 6 two typical time series of price returns, together with a chosen threshold Q = -0.037, which corresponds to an average inter-occurrence time $R_Q = 70$ [106,107]. The quantity R_Q monotonically increases with |-Q| in each one of the examples shown in Figures 7 and 8. For a fixed value of R_Q , we verify that $p_Q(r) \propto e_{q \text{ threshold } r}^{-\beta \text{threshold } r}$, with:

$$q_{threshold} = 1 + q_0 \ln(R_O/2) \quad (q_0 \simeq 0.168).$$
 (22)

See the illustrations in Figures 9 and 10. The fact that we have analytically $p_Q(r)$ enables us to straightforwardly obtain an explicit expression for the risk function $W_Q(t; \Delta t)$, which is defined as the probability of having once again a fluctuation larger than |-Q| within an interval Δt at time *t* after the last large fluctuation. It can be shown [106,109] that:

$$W_Q(t;\Delta t) \equiv \frac{\int_t^{t+\Delta t} p_Q(r)dr}{\int_t^{\infty} p_Q(r)dr} = 1 - \left[1 + \frac{\beta_{threshold}(q_{threshold}-1)\Delta t}{1+\beta_{threshold}(q_{threshold}-1)}\right]^{\frac{q_{threshold}-2}{q_{threshold}-1}} = 1 - \frac{e_{\bar{q}}^{-(\beta_{threshold}/\bar{q})(t+\Delta t)}}{e_{\bar{q}}^{-(\beta_{threshold}/\bar{q})t}}$$
(23)

with $\tilde{q} \equiv 1/(2 - q_{threshold})$. See Figure 11.



Figure 6. Illustration of the relative daily price returns X_i of the IBM stock between (**a**) January 2000 and June 2010 and (**b**) 27 August and 23 October 2002. The red line shows the threshold $Q \simeq -0.037$, which corresponds to an average inter-occurrence time of $R_Q = 70$. In (b), the inter-occurrence times are indicated by arrows. From [106].



Figure 7. The mean inter-occurrence time R_Q vs. the absolute value of the loss threshold -Q. The continuous curves are fittings with $R_Q = A e_{q_{inter}}^{B_{inter}|Q|} = A[1 + (1 - q_{inter})B_{inter}|Q|]^{1/(1-q_{inter})}$. Top left: For the exchange rate of the U.S. Dollar against the British Pound, the index S&P500, the IBM stock and crude oil (West Texas Intermediate (WTI)), from left to right in the plot; the corresponding values for q_{inter} are 0.95, 0.92, 0.97, 0.927 (with A = 2, 2.04, 1.95, 2.02 and $B_{inter} = 240$, 175, 95, 60). Similarly for the top right, bottom left and bottom right plots. From [107].



Figure 8. The mean inter-occurrence time R_Q versus the absolute value of the loss threshold -Q: $\ln_{q_{inter}}(R_Q/A)$ versus the $B_{inter}|-Q|$ representation of the same data of Figure 7. The continuous curve is a fitting with $R_Q = Ae_{q_{inter}}^{B_{inter}}|-Q|$. From [107].



Figure 9. (a) The distribution function of the inter-occurrence times for the relative daily price returns X_i of IBM in the period 1962–2010. The data points belong to $R_Q = 2, 5, 10, 30$ and 70 (in units of days), from bottom to top. The full lines show the fitted *q*-exponentials $p_Q(r) \propto e_{qhtreshold}^{-\beta_{threshold}}$ for typical values of R_Q . (b) The dependence of the parameters $\beta_{threshold}$ (squares, lower curve) and $q_{threshold}$ (circles, upper curve) on R_Q in the $q_{threshold}$ -exponential. (c) Confirmation that, for $R_Q = 2$, the distribution function is a simple exponential (i.e., $q_{threshold} = 1$). The straight line is proportional to 2^{-r} . From [106].



Figure 10. The distribution function of the inter-occurrence times (as in Figure 9a) for the relative daily price returns of 16 examples of financial data, taken from different asset classes (stocks, indices, currencies, commodities). The assets are: (i) the stocks of IBM, Boeing (BA), General Electric (GE), Coca-Cola (KO); (ii) the indices Dow Jones (DJI), Financial Times Stock Exchange 100 (FTSE), NASDAQ, S&P 500; (iii) the commodities Brent Crude Oil, West Texas Intermediate (WTI), Amsterdam-Rotterdam-Antwerp gasoline (ARA), Singapore gasoline (SING); and (iv) the exchange rates of the following currencies versus the U.S. Dollar: Danish Crone (DKK), British Pound (GBP), Yen, Swiss Francs (SWF). The full lines show the fitted *q*-exponentials, which are the same as in Figure 9a. From [106].



Figure 11. Universal risk function $W_Q(t, \Delta t)$ from Equation (23) for the inter-occurrence time $R_Q = 100$ and for the intervals $\Delta t = 1, 5, 20, 100$ days (from bottom to top). From [106].

2.4. Wealth

Wealth inequality within a given country is a classical and most important matter, which can be characterized within *q*-statistics as shown in [105]: see Figures 12 and 13. The larger the index $q_{inequality}$ is, the larger the inequality. As we verify, the U.K. and Germany are more egalitarian countries than the U.S. and Brazil. In addition to that, inequality appears to increase in the U.S. and Brazil, at least during the years indicated in the plots.



Figure 12. Binned inverse cumulative distribution of the county, PI/PI_0 (U.S.) and GDP/GDP_0 (Brazil, Germany and U.K.), where PI denotes the Personal Income and GDP denotes the Gross Domestic Product of countries. Three distributions are displayed for comparison: (i) *q*-Gaussian (with $\beta_{q'} = 0$) (dot-dashed); (ii) (*q*; *q'*)-Gaussian (solid) and (iii) log-normal (dashed lines). (**a**,**b**) present insets with a linear-linear scale, to make more evident the quality of the fitting at the low region (in (**c**,**d**), the (*q*; *q'*)-Gaussian and the log-normal curves are superposed and, so, are visually indistinguishable). The positions of the knees are indicated. The ankle is particularly pronounced in (c), though it is also present in the other cases. From [105], where further details are available.

Another index that characterizes the wealth of a country and its inequalities is associated with the prices of the land. See in Figure 14 (from [105]) an illustration for Japan, where $q_{jand \ price} = 2.136$.



Figure 13. Evolution of parameter *q* for the U.S. (squares), Brazil (circles), the U.K. (up triangles) and Germany (down triangles). The parameters q' (for each country) are constant for all years: $q'_{Brazil} = 2.1$, $q'_{USA} = 1.7$, $q'_{Germany} = 1.5$, $q'_{UK} = 1.4$. Lines are only guides to the eyes. As we verify, in some cases, the index *q* remains invariant along time, whereas in others, it evolves; the functional forms remain however the same as indicated in Figure 12. From [105].



Figure 14. Inverse cumulative probability distribution of Japanese land prices for the year 1998. The solid curve is a *q*-Gaussian with q = 2.136, which corresponds to the slope -1.76, and $1/\sqrt{\beta_q} = 188,982$ Yen. From [105], where further details are available.

3. Conclusions and Perspectives

We have described a variety of financial and economic properties with a plethora of q-indices, such as q_{return} , q_{volume} , $q_{volatility}$, q_{inter} , $q_{threshold}$, q_0 , \tilde{q} , $q_{inequality}$, $q_{land price}$. For a given system, how many independent indices should we expect? The full answer to this question remains up to now elusive. It seems however that only a few of them are essentially independent, all of the others being (possibly simple) functions of those few. Such an algebraic structure was first advanced and described in [119] and has been successfully verified in the solar wind [53] (see also [6] and the references therein) and elsewhere; it has recently been generalized [120,121] and related to the Moebius group. The central elements of these algebraic structures appear to constitute what is currently referred to in the literature as q-triplets [122]. The clarification and possible verification of such structures constitutes nowadays an important open question, whose further study would surely be most useful.

Another crucial question concerns the analytic calculation from first principles of some or all of the above q-indices. This is in principle possible (as illustrated in [63–66]), but it demands the complete

knowledge of the microscopic model of the specific class of the complex system. For the full set of the *q*-indices shown in the present overview, such models are not available, even if they would be very welcome.

Let us finally emphasize that many other statistical approaches exist for the quantities focused on in the present overview. However, as announced in the title of this paper, this is out of the present scope. The present paper is one among various others belonging to the same Special Issue of the journal Entropy. The entire set of articles is expected to enable comparisons between these many approaches.

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